Estimating Identifiable Causal Effects through Double Machine Learning

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Learning causal effect: 2-step procedures





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Represent the causal query Q = P(y|do(x)) as a function of P and Step 1. set of causal assumption A (i.e., Q = F(P, A)). (Identification, ID)











Learning causal effect: 2-step procedures

Represent the causal query Q = P(y|do(x)) as a function of P and Step 1. set of causal assumption A (i.e., Q = F(P, A)). (Identification, ID)

Estimate the identified estimand Q = F(P, A) from finite samples D. Step 2. (Estimation)















Strength Completeness of Identification algorithm

 There exist *sound* and *complete* identification algorithms for determining whether a causal query **Q** can be represented as a functional of P (i.e., Q=F(P, A)) from a given causal graph.





Ζ

Χ There exist sound and c identification algorithms **Backdoor graph** whether a causal query G can be represented as a functional of P (i.e., Q=F(P, A)) from a given causal graph.



Weakness **No sample-efficient estimator**

 Estimation has been mainly done on backdoor/ignorability assumption.



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Weakness **No sample-efficient estimator**

- Estimation has been mainly done on backdoor/ignorability assumption.
- For general identifiable estimands, it's not known (neither obvious) how to estimate the causal effect sample-efficiently.















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- **Observation** Causal functional P is a arithmetic of two adjustments:

P(x, y | r, w) adjusting over W=w

 $P(x \mid r, w)$ adjusting over W=w





A causal functional is expressible as a function of adjustments. Then, a BD estimator might be applicable in estimating such functional.

 $P(y \mid do(x)) = \frac{\sum_{w} P(x, y \mid r, w) P(w)}{\sum_{w} P(x \mid r, w) P(w)}$

ntuition

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A causal functional is expressible as a function of adjustments. Then, a BD estimator might be applicable in estimating such functional.

> **Observation** — Causal functional P is a

What BD estimators should we leverage?





Recap: Classic BD estimators $Q = \sum P(y | x, z) P(z)$



Backdoor graph

Estimand (Q)

Estimator (\widehat{Q})

For correct estimation

For \sqrt{N} -consistency





Recap: Classic BD estimators $Q = \sum P(y | x, z) P(z)$ **Backdoor graph Inverse Probability Weight**

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Inverse Probability Weight





BD	estimators	

 $Q = \sum P(y | x, z) P(z)$

Inverse Probability Weight





C BD estimators $Q = \sum_{z} P(y x, z) P(z)$				
ability Weight	Regression			
$-I_y(Y_{(i)})$				





BD estimators				
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	·			





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$-I_y(Y_{(i)})$	$\frac{1}{N} \sum_{i=1}^{N} \widehat{P}(y \mid x, Z_{(i)})$			





















 $Q = \sum P(y | x, z) P(z)$

Risk of classic estimators

Incorrectly estimated if estimates of nuisances are misspecified.

Not \sqrt{N} -consistent if estimates of nuisances converges slower









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Model misspecification is common in practice when the data generating process is complicated \Rightarrow Incorrect estimation.

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Backdoor graph



-consistency.

than $o_P(N^{-1/2})$

Estimand (*Q*)

Estimator (Q)

For correct estimation

For \sqrt{N} -consistency

 $Q = \sum P(y | x, z) P(z)$

Risk of classic estimators

- Model misspecification is common in practice when the data generating process is complicated \Rightarrow Incorrect estimation.
- Slow convergence is common for flexible ML models \Rightarrow Not \sqrt{N}
- **Incorrectly estimated** if estimates of nuisances are **misspecified**.
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Recap: <u>Classic BD estimators</u>

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Estimator (\widehat{Q})

For correct estimation

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I NEED SOMETHING ROBUST

tors

Not \sqrt{N} -consistent if estimates of nuisances converges slower









Backdoor graph

Double Machine Learning Estimator

Estimand

Representation of ${\boldsymbol{\mathcal{Q}}}$

Estimators

Estimator \widehat{Q}

For correct estimation

For \sqrt{N} -consistency

$Q = \sum_{z} P(y | x, z) P(z)$





Backdoor graph

Double Machine Learning Estimator

Estimand

Representation of Q

Estimators

Estimator \widehat{Q}

For correct estimation

For \sqrt{N} -consistency

 $\mathbb{E}\left[\frac{I_x}{P(X)}\right]$

$$Q = \sum_{z} P(y | x, z) P(z)$$

$$\frac{P(X)}{X|Z} \left(I_{y}(Y) - P(y|x,Z) \right) + P(y|x,Z) \right]$$





Backdoor graph

Double Machine Learning Estimator

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Representation of Q

Estimators

Estimator Q

For correct estimation

For \sqrt{N} -consistency





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$$I_y(Y) - \widehat{P}(y|x,Z) + \widehat{P}(y|x,Z)$$





Backdoor graph

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 $\frac{1}{N}\sum_{i=1}^{N} \frac{I_{x}(X_{(i)})}{\widehat{P}(X_{(i)}|Z_{(i)})} \left(I_{y}(Y) - \widehat{P}(y|x,Z)\right) + \widehat{P}(y|x,Z)$ where training and evaluation of \widehat{P} are done with two distinct sets of samples ("**Cross fitting**")







Backdoor graph

Double Machine Learning Estimator

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Representation of Q

Estimators

Estimator Q

For correct estimation

("Doubly robustness")

For \sqrt{N} -consistency



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("Debiasedness") Estimates for nuisance converges at $O_P(N^{-1/4})$.




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Estimator Q

For correct estimation

("Doubly robustness")

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DML estimator is robust!

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- nuisance converges slowly.

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Estimand

Representation of Q

Estimators Estimator Q

В

- nuisance converges slowly.

Is DML estimators applicable for estimating any identifiable causal functional?

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Neyman orthogonal score ϕ

For the target estimand ψ (e.g., $\psi = P(y | do(x)))$ and nuisances η (e.g., $\eta = \{P(y | x, z), P(x | z)\})$, a function $\phi(\mathbf{V}; \psi, \eta)$ is called a Neyman orthogonal score if

- 1. (Moment condition) $\mathbb{E}_P \left| \phi(\mathbf{V}; \psi, \eta_0) \right| = 0$ where η_0 is a true nuisance, and
- 2. (Orthogonality) $(\partial/\partial\eta)\mathbb{E}_P \left[\phi(\mathbf{V};\psi,\eta)\right]|_{\eta=\eta_0} = 0.$





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DML estimator

An estimator for ψ that is composed of $\hat{\eta}$, such that

- 1. (Neyman orthogonal score) the estimand is based on Neyman orthogonal score; and
- 2. (Cross-fitting) training and evaluating $\hat{\eta}$ is done with two distinct sets of samples.







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Debiasedness

- Without any smoothness & complexity constraints on the model for $\hat{\eta}$, DML estimator achieves debiasedness property:
 - (**Debiasedness**) \sqrt{N} -consistent whenever $\hat{\eta}$ converges to true nuisance at $N^{-1/4}$ rate.

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Research question

Construct a DML estimator \hat{Q} that is

- robust against model misspecification (doubly robust) and slow convergence (debiased); and
- working for any identifiable causal functional. (Complete)





Statistical p

Doubly robustness

Statistical property		Causal property	
Doubly robustness	Debiasedness	Beyond BD	Any identifiable











Statistical property

Doubly robustness

Jung, Tian, Bareinboim (2020a)

Fulcher et al (2019), Bhattacharya et al (2020)







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DML-ID (Jung, Tian, Bareinboim, 2021)









Result 0. (mSBD)



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(X, Y) (called multi-outcome sequential back-door (mSBD)).

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• Represent any identifiable causal estimand Q=F(P) as an arithmetic (multiplication,





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Result 1. (DML-ID)

- Represent any identifiable causal estimand Q=F(P) as an arithmetic (multiplication, ratio, marginalization) of mSBDs.
- Derive the Neyman orthogonal score for any arbitrary causal functional in a form of an arithmetic of scores of mSBDs.







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Construct a DML estimator working for any identifiable causal functional.







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(Informal) Multi-outcome **Sequential Back-door (mSBD)**

 $\mathbf{Z} = \{\mathbf{Z}_1, \dots, \mathbf{Z}_n\}$ satisfies *mSBD* criterion relative to $\{\mathbf{X}, \mathbf{Y}\}$ if non-causal path b/w $X_i \in \mathbf{X}$ and $\mathbf{Y}_i \in \mathbf{Y}$ are blocked by \mathbf{Z}_i conditioned on { $\mathbf{X}^{(i-1)}, \mathbf{Y}^{(i-1)}, \mathbf{Z}^{(i-1)}$ }.

(Implication): No unmeasured confounders b/w X and Y.



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mSBD adjustment

 $P(\mathbf{y} | do(\mathbf{x})) = \sum P(\mathbf{v}_i | \mathbf{v}^{(i-1)}),$ $z \quad V_i \in \{Y, Z\}$

• for $\mathbf{V}_i = \mathbf{Y}_i$, $\mathbf{V}^{(i-1)} = {\mathbf{X}^{(i)}, \mathbf{Z}^{(i)}, \mathbf{Y}^{(i-1)}};$ and • for $\mathbf{V}_i = \mathbf{Z}_i, \mathbf{V}^{(i-1)} = \{\mathbf{Y}^{(i-1)}, \mathbf{X}^{(i-1)}, \mathbf{Z}^{(i-1)}\}$.



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(Implication): No unmeasured confounders $b/w \mathbf{X}$ and \mathbf{Y} .



mSBD adjustment

 $P(\mathbf{y} | do(\mathbf{x})) = \sum P(\mathbf{v}_i | \mathbf{v}^{(i-1)}),$ $z V_i \in \{Y, Z\}$ • for $\mathbf{V}_i = \mathbf{Y}_i, \mathbf{V}^{(i-1)} = {\mathbf{X}^{(i)}, \mathbf{Z}^{(i)}, \mathbf{Y}^{(i-1)}};$ and

• for $\mathbf{V}_i = \mathbf{Z}_i, \mathbf{V}^{(i-1)} = \{\mathbf{Y}^{(i-1)}, \mathbf{X}^{(i-1)}, \mathbf{Z}^{(i-1)}\}$.







Neyman orthogonal score for mSBD

 $\phi(\mathbf{V}; \psi, \eta) = \sum_{i=1}^{n} W_i (H_{i+1} - H_i), \text{ where}$ $H_i = P_{\mathbf{X}}(\mathbf{y}^{\geq i-1} | \mathbf{Z}^{(i-1)}, \mathbf{y}^{(i-2)}) \text{ and}$ $W_i = \prod_{p=1}^{i} \frac{I_{x_p}(X_p)}{P(x_p | \mathbf{Z}^{(p)}, \mathbf{x}^{(p-1)}, \mathbf{y}^{(p-1)})}$



Neyman orthogonal score for mSBD

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Our Approach for marrying DML + Identification (DML-ID)

Result 0. (mSBD)

 Derive the Neyman orthogonal score when there are no unmeasured confounders b/w (X, Y) (called multi-outcome sequential back-door (mSBD)).

Result 1. (DML-ID)

- Represent any identifiable causal estimand Q=F(P) as an arithmetic (multiplication, ratio, marginalization) of mSBDs.
- Derive the Neyman orthogonal score for any arbitrary causal functional in a form of an arithmetic of scores of mSBDs.

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Construct a DML estimator working for any identifiable causal functional.







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Complete ID algorithm (Tian and Pearl, 2003).

• A causal functional $P(\mathbf{y} | do(\mathbf{x}))$ is identifiable if and only if it is represented as an arithmetic of C-factors of C-component C_i in G (Tian and Pearl, 2003, Huang and Valtorta 2006).





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Q = P(y|do(x))G ID algorith





 $\sum_{w} Q[W, X, Y] \qquad \sum_{w} P(x, y \mid r, w) P(w)$ $= \overline{\sum_{w,y} Q[W, X, Y]} - \overline{\sum_{w} P(x \mid r, w) P(w)}$















 $\sum_{w} P(x, y \mid r, w) P(w)$ $\sum_{W} Q[W, X, Y]$ $\overline{\sum_{w,y} Q}[W, X, Y]$ $\sum_{w} P(x \mid \overline{r, w}) P(w)$











Result 1. Derivation of Neyman Orthogonal Score



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Given representation of $\mathbf{Q} = \mathbf{A}(M[\mathbf{C_1}], M[\mathbf{C_2}], \dots, M[\mathbf{C_d}])$, a Neyman orthogonal score is given as



- Thm. 2: Derivation of Neyman Orthogonal score

$\sum_{i=1}^{n} \phi_{M_{i}} \frac{\partial}{\partial M_{i}} A\left(M[\mathbf{C}_{1}], \cdots, M[\mathbf{C}_{d}]\right)$

Neyman orthogonal score for the mSBD adjustment M_i=M[**C**_i]



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Result 2 — Empirical evidence



 $P(y \mid do(x)) = \frac{\sum_{w} P(x, y \mid r, w) P(w)}{\sum_{w} P(x \mid r, w) P(w)}$




$P(y \mid do(x)) = -$

$$\frac{\sum_{w} P(x, y \mid r, w) P(w)}{\sum_{w} P(x \mid r, w) P(w)}$$

We compare this estimator with **Plug-in estimator**, which estimates cond.prob $\{P(x, y | r, w), P(x | r, w), P(w)\}$ and plugs those in for computing the estimand.





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- arbitrary causal functional.

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 A reason that we compare with the plug-in estimator is because this is only viable estimator that is working for



Debiasedness

Doubly robustness



Debiasedness

 $\widehat{P}(x, y | r, w), \ \widehat{P}(r | w) \text{ converges to}$ true P(x, y | r, w), P(r | w) at a rate $N^{-1/4}$ **Doubly robustness**





Doubly robustness





Doubly robustness

(x, y r, w) misspecified.	$P(r \mid w)$ misspecified.
	<u>.</u>







Conclusion



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orthogonal score for estimands of any identifiable causal effects.



• **Result 1** — We develop a systematic procedure for deriving Neyman



Conclusion

- orthogonal score for estimands of any identifiable causal effects.
- bias.



Result 1 — We develop a systematic procedure for deriving Neyman

Result 2 — We develop DML estimators for any identifiable causal effect, which enjoy debiasedness and doubly robustness against model misspecification and



Thank you for listening!